Unconfounded Treatment Assignment

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- Move toward the setting of observational studies.
- Relax the classical randomized experiment assumption.

Assignment Assumption

- Probabilistic assignment
- Individualistic assignment
- Unconfounded assignment
 - Pr(W|X, Y(0), Y(1)) are free from dependence on the potential outcomes.
 - In combination with individualistic assignment,

$$Pr(W|X, Y(0), Y(1)) = c \cdot \prod_{i=1}^{N} e(X_i)^{W_i} (1 - e(X_i))^{1 - W_i}$$

where e(x) is propensity score.

• The assumption is extremely widely used.

$$Y_i^{obs} = \alpha + \tau_{sp} \cdot W_i + X_i\beta + \epsilon_i$$

• It is assumed that $\epsilon_i \perp W_i, X_i$

• In the potential outcome formulation, we have

$$Y_i(0) = \alpha + X_i\beta + \epsilon_i$$

 $Y_i(1) = Y_i(0) + \tau_{sp}$

• Then ϵ_i is a function of $Y_i(0)$ and X_i given the parameters

$$Pr(W_i = 1 | X_i, Y_i(0), Y_i(1)) = Pr(W_i | \epsilon_i, X_i) = Pr(W_i | X_i)$$

Why Is Unconfoundedness an Important Assumption?

- By assuming unconfoundedness we can compare the particular treated units with control units.
- If there is an alternative of unconfoundedness, it must involve looking for a comparison which is different in terms of observed pre-treatment variables.
- In many cases it would appear implausible.

• Assuming individualistic assignment and unconfounded assignment,

$$Pr(W|X, Y(0), Y(1)) = c \cdot \prod_{i=1}^{N} e(X_i)^{W_i} (1 - e(X_i))^{1 - W_i}$$

• A balancing score b(x) is a function of the covariates such that

 $W_i \perp X_i | b(X_i)$

• The probability of receiving the treatment given the covariates is free of dependence on the covariates given the balancing score.

Lemma

The propensity score is a balancing score.

Lemma

The propensity score is the coarsest balancing score. That is, the propensity score is a function of every balancing score.

- Prior to estimating causal effects it is important to conduct design phase of an observational study.
 - Assessing Balance CH13, CH14
 - Subsample selection using matching on the propensity score CH15
 - Subsample selection through trimming using propensity score CH16

- Discuss five broad classes of strategies for estimation
 - Model-Based imputation
 - Regression estimators
 - Weighting estimators that use the propensity score
 - Blocking estimators that use the propensity score (CH17)
 - Matching Estimators (CH18)
- Blocking and matching are relatively attractive because of the robustness properties that stem form the combination of methods.

• Weighting exploits the two equalties

$$\mathbb{E}\left[\frac{Y_i^{obs} \cdot W_i}{e(X_i)}\right] = \mathbb{E}_{sp}[Y_i(1)]$$
$$\mathbb{E}\left[\frac{Y_i^{obs} \cdot (1 - W_i)}{1 - e(X_i)}\right] = \mathbb{E}_{sp}[Y_i(0)]$$

• Horvitz and Thompson introduced estimator of ATE as

$$\hat{\tau}^{ht} = \frac{1}{N} \sum_{i:W_i=1} \lambda_i \cdot Y_i^{obs} - \frac{1}{N} \sum_{i:W_i=0} \lambda_i \cdot Y_i^{obs}$$

where

$$\lambda_i = \frac{1}{e(X_i)^{W_i} \cdot (1 - e(X_i))^{1 - W_i}}$$

- Let b_j , $j = 0, 1, \dots, J$ denote the subclass boundaries, with $b_0 = 0$, $b_J = 1$.
- Let B_i(j) be a binary indicator, equal to 1 if b_{j-1} < e(X_i) < b_j and zero otherwise.

• Estimate the finite-sample average effect in subclass j by

$$\hat{ au}^{dif}(j) = rac{\sum_{i:B_i(j)=1}Y_i \cdot W_i}{\sum_{i:B_i(j)=1}W_i} - rac{\sum_{i:B_i(j)=1}Y_i \cdot (1-W_i)}{\sum_{i:B_i(j)=1}(1-W_i)}$$

· Estimate the overall finite-sample average effect of the treatment by

$$\hat{\tau}^{strat} = \sum_{j=1}^{J} \frac{N(j)}{N} \cdot \hat{\tau}^{dif}(j)$$

where $N(j) = \sum_{i=1}^{N} B_i(j)$

• For a given treated unit with a particular set of values for the covariates, one looks for a control unit with as similar a set of covariates as possible.

Efficiency Bounds

• Define

$$\mu_{c}(x) = \mathbb{E}_{sp}[Y_{i}(0)|X_{i} = x], \ \mu_{t}(x) = \mathbb{E}_{sp}[Y_{i}(1)|X_{i} = x]$$
$$\sigma_{c}^{2}(x) = \mathbb{V}_{sp}[Y_{i}(0)|X_{i} = x], \text{ and } \sigma_{t}^{2}(x) = \mathbb{V}_{sp}[Y_{i}(1)|X_{i} = x]$$

• The super-population average treatment effect defined as

$$\tau_{sp} = \mathbb{E}_{sp}[Y_i(0) - Y_i(1)] = \mathbb{E}_{sp}[\tau_{sp}(X_i)]$$

where

$$\tau_{sp}(x) = \mu_t(x) - \mu_c(x) = \mathbb{E}_{sp}[Y_i(1) - Y_i(0)|X_i = x]$$

• The finite-sample average effect conditional on the values of the pre-treatment variables is defined as

$$\tau_{cond} = \frac{1}{N} \sum_{i=1}^{N} \tau_{sp}(X_i)$$

• Under unconfoundedness and probabilistic assignment, and without additional functional form restrictions beyond smoothness, the sampling variance bound for estimators for τ_{sp} , normalized by the sample size is

$$\mathbb{V}_{sp}^{eff} = \mathbb{E}_{sp} \left[\frac{\sigma_c^2(X_i)}{1 - e(X_i)} + \frac{\sigma_t^2(X_i)}{e(X_i)} + \left(\tau_{sp}(X_i) - \tau_{sp} \right)^2 \right]$$

• The sampling variance bound for estimators for τ_{cond} is

$$\mathbb{V}_{cond}^{eff} = \mathbb{E}_{sp} \left[\frac{\sigma_c^2(X_i)}{1 - e(X_i)} + \frac{\sigma_t^2(X_i)}{e(X_i)} \right]$$